

Adaptive Partition Factor Analysis

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Joint Work with Elena Bortolato, UPF, Barcelona



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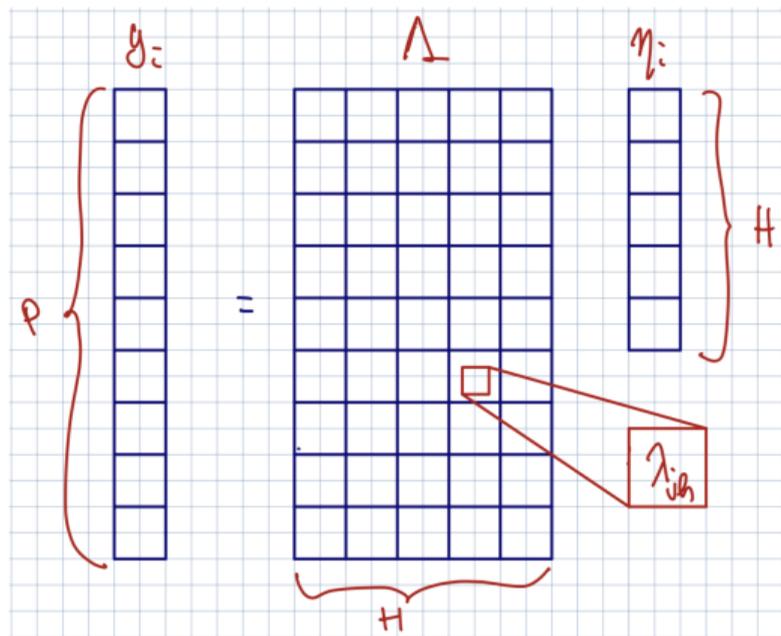
- we observe a p -dimensional vector of item responses (binary / ordinal) to p psychological questionnaire items
- observations are grouped in S independent studies (or cohorts), with $s = 1, \dots, S$
- in each study we have multiple participants, for a total of n individuals overall ($i = 1, \dots, n$)
- data y_{is} is the p -dimensional response vector for individual i in study s
- goal: infer latent psychological dimensions (factors) that are comparable across studies

- we consider this data in a “multi-study” fashion
- other examples:
 - medical studies performed in different hospitals
 - genomics studies performed with different technological platforms
 - ...
- we are going to follow a factor model approach, and specifically a **Multi-Study Factor Analysis** (MSFA, De Vito et al. 2017) approach
- MSFA extensions and generalizations:
 - Roy et al. (2021) proposed a perturbed factor analysis that focuses on inferring the shared structure while making use of subject-specific perturbations
 - Grabski et al. (2023) proposed a model allowing for partially-shared latent factors
 - Chandra et al. (2024) proposed a class of subspace factor models with appealing identifiability properties

Factor Analysis (FA)

$$y_i = \Lambda \eta_i + \epsilon_i$$

- y_i : i -th p -variate random variable;
- Λ : $p \times H$ factor loadings matrix;
- η_i : i -th vector of H latent factors.



- Most of the interest of FA revolves around the concept of **interpretability**;
- Interpretation of factor models is assigning a meaning to the latent factors and then to their impact on the observed data;
- this is promoted by the concept of **sparsity** in many ways
- In this talk I will exploit sparsity within a MSFA framework

- Multi-study factor analysis (MSFA) assumes the existence of both shared latent factors and study-specific latent factors
- Specifically

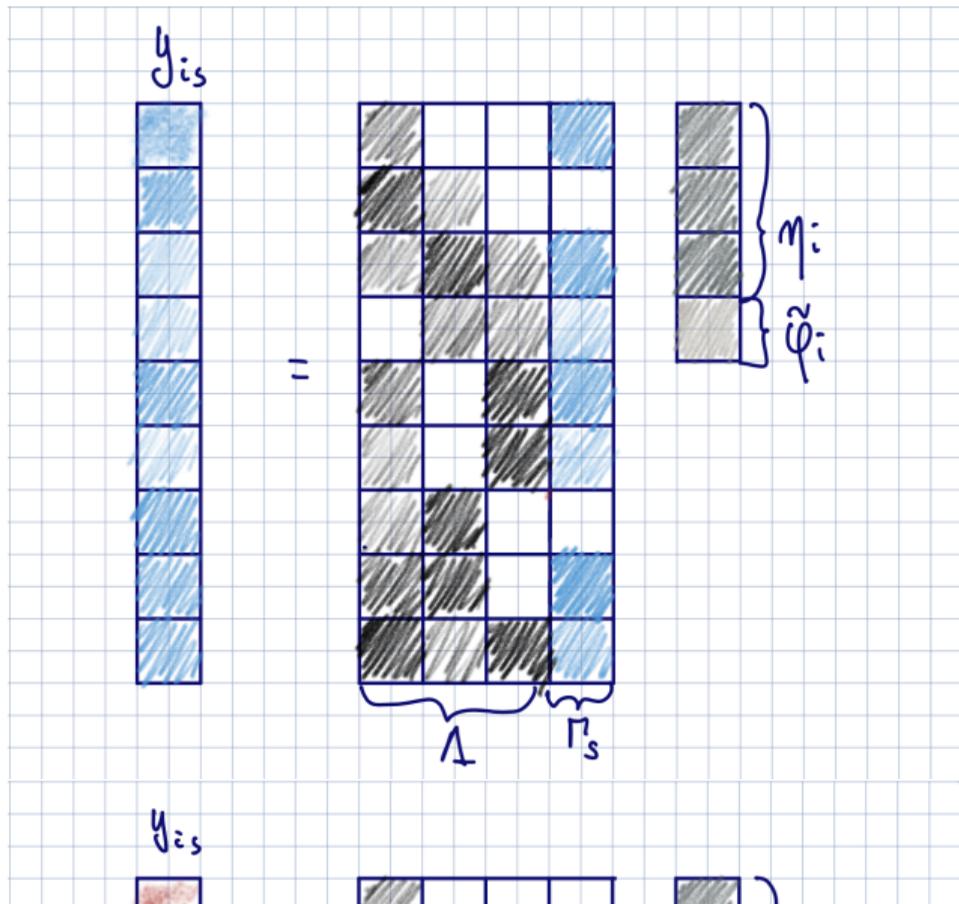
$$y_{is} = \Lambda \eta_{is} + \Gamma_s \varphi_{is} + \epsilon_{is} \quad (1)$$

where Γ_s is a (study-specific) factor loading matrix of dimension $p \times k_s$, with $k_s \ll p$ possibly different in each study, and φ_{is} its corresponding latent factor.

- The resulting marginal distribution of y_{is} is Gaussian with covariance

$$\Omega_s = \Lambda \Lambda^T + \Gamma_s \Gamma_s^T + \Sigma_s.$$

MSFA graphically

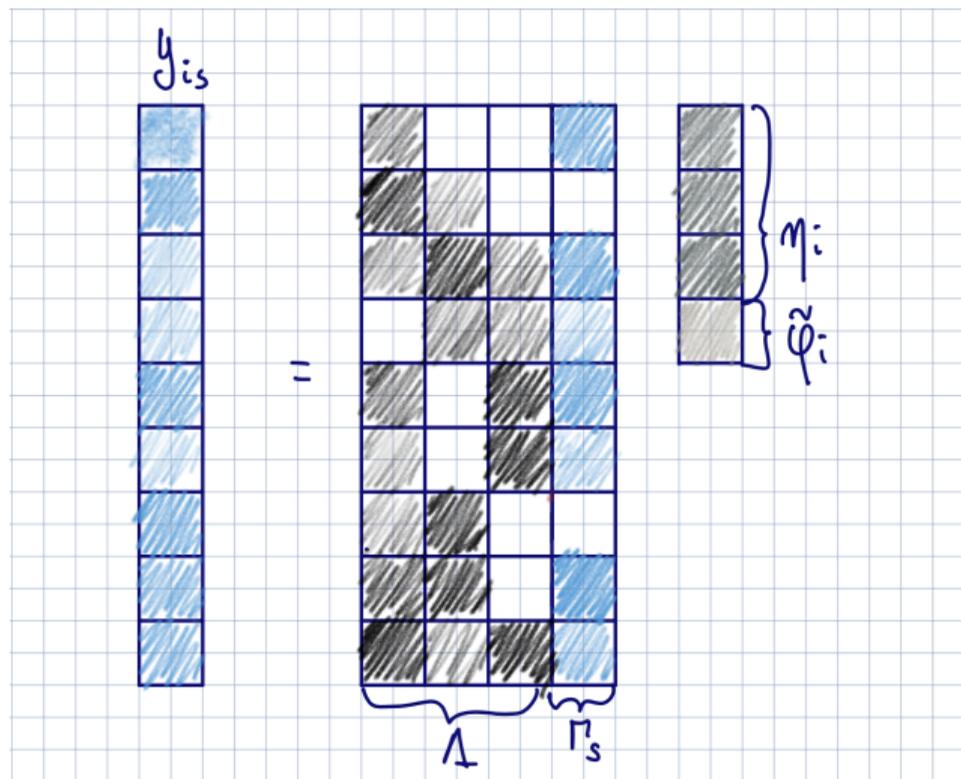


- Rewrite the MSFA as

$$y_i = \Lambda \eta_i + \Gamma \varphi_i + \epsilon_i. \quad (2)$$

- Here $\Gamma = (\Gamma_1, \dots, \Gamma_S)$ concatenate along columns all the study-specific factor loading matrices into a $p \times k$ matrix with $k = \sum_{s=1}^S k_s$
- φ_i is a k -dimensional augmented vector containing the original φ_{is} framed with suitable pattern of zeroes.

Rethinking the MSFA graphically



Fixed partition vs Adaptive partition

- MSFA permits precisely S study-specific loading matrices Γ_s
- practical scenarios often present more complex situations:
 - two or more studies may present high homogeneity, potentially sharing identical or nearly identical latent representations
 - some studies may involve a highly heterogeneous group of subjects, possibly leading to two or more sub-populations displaying distinct latent representations
- An **adaptive** partition that accounts for the above situations (and beyond) would be useful!

- As customary in factor analysis marginalize out the latent factors $\eta_i \sim N(0, I_p)$,

$$y_i \sim N(\Gamma \varphi_i, \Lambda \Lambda^T + \Sigma). \quad (3)$$

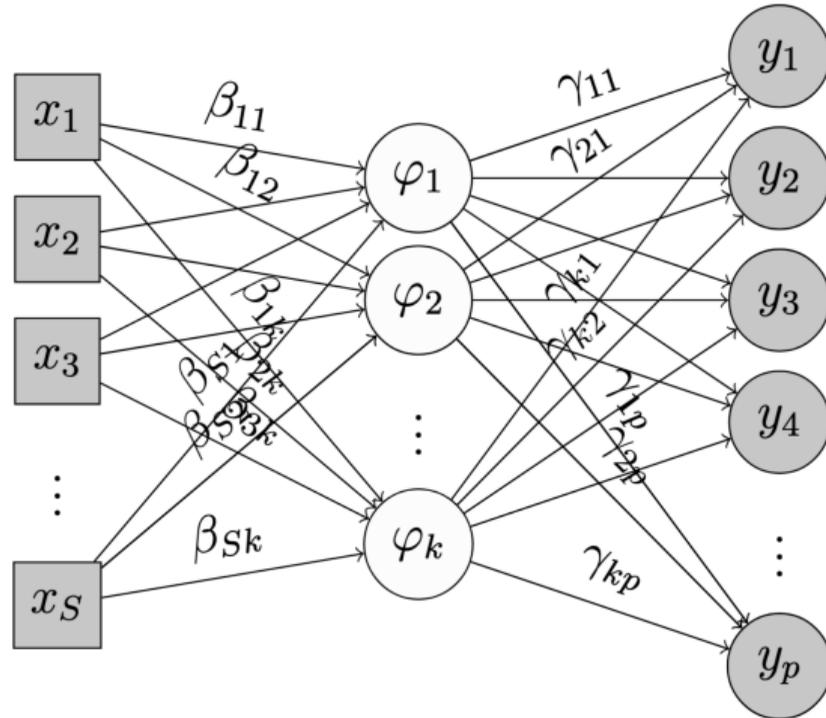
- Let x_i be a dummy variable scharacterizing the study to which unit i belongs
- The h -th element of the vector φ_i is $\varphi_{ih} = \tilde{\varphi}_{ih} \psi_{ih}$ where $\tilde{\varphi}_{ih}$ is a continuous random variable and $\psi_{ih} = f_h(x_i)$ with f_h a deterministic activation function.
- for MSFA $f_h(x_i) = x_i^T \mathbf{1}_S$ where $\mathbf{1}_S$ is a S dimensional vector of ones

- What about incorporating the information contained in the x_i 's in a more flexible manner?
- For example with

$$\psi_{ih} = f_h(\mathbf{x}_i^T \beta_h). \quad (4)$$

- Which leads to a specific single layer neural network where
 - the x_i s are the input variables,
 - y_i are output variables,
 - φ_i s are the nodes of the hidden layer
 - f_h are the activation functions
 - β_h are the weights between the input and the hidden layer,
 - Γ are the weights between the hidden layer and the outcome

From factor analysis to neural networks



- We encode the study-specific membership into categorical the variables x_i
- The study-specific latent factors are assumed

$$\varphi_{ih} \sim N(\mathbf{0}, \psi_{ih}(x_i)\tau_h).$$

- We then assume a dependence between the scale parameters of the group-specific latent factors and the group indicator x_i , as follows:

$$\psi_{ih}(x_i) \sim \text{Ber}\{\text{logit}^{-1}(x_i^\top \beta)\}.$$

- The shrinkage prior on the elements $\varphi_{ih}(x_i)$ enables and promotes, yet does not mandate, the sparse representation of MSFA
- Wide range of scenarios including:
 - two or more studies exhibit high homogeneity and share nearly identical latent representations,
 - some studies involve highly heterogeneous groups of subjects, potentially resulting in two or more sub-populations with distinct latent structures,
 - any combinations of the above.

Time for discussion?

Other (more technical) stuff...

- 1 Introduction
- 2 From MSFA to APAFA
- 3 Theoretical properties
- 4 Empirical performance
 - Simulation Study
 - Illustration: Finnish birds co-occurrence data
- 5 References

- Factor loadings are not identifiable due to rotations, i.e .

$$\Lambda\Lambda^T = \tilde{\Lambda}\tilde{\Lambda}^T, \quad \text{if } \tilde{\Lambda} = \Lambda P, \quad \text{and } PP^T = I_k$$

- In the multi-study context and, specifically, in

$$y_i \sim N(0, \Omega_i), \quad \Omega_i = \Lambda\Lambda^T + \Gamma \text{diag}\{\psi_i\}\Gamma^T + \Sigma$$

we may additionally face “information switching”

Definition (Information Switching)

Let S_n the number of distinct groups, i.e. $S_n = |\cup_{i=1}^n \Omega_i|$. Denote with Ψ the $n \times k$ matrix that stacks in distinct rows all $\psi_i = (\psi_{i1}, \dots, \psi_{ik})$ and with Ψ_h its generic column with, $\Psi_h \neq \mathbf{1}_n$ for all $h = 1, \dots, k$. Let Ω_s^* and ψ_s^* ($s = 1, \dots, S_n$) be the distinct values of Ω_i and ψ_i , respectively. Similarly, let $W_s^* = \Omega_s^* - \Lambda\Lambda^\top - \Sigma = \Gamma \text{diag}\{\psi_s^*\} \Gamma^\top$. The model suffers from information switching if there exist $\tilde{\Gamma} \neq \Gamma$ and $\tilde{\Psi} \neq \Psi$ such that $W_s^* = \tilde{\Gamma} \text{diag}\{\tilde{\psi}_s^*\} \tilde{\Gamma}^\top$ for all s , with $\tilde{\Psi}_h = \mathbf{1}_n$ for at least one h .

Theorem

If $\Psi_h \neq \mathbf{1}_n$ for all $h \in \{1, \dots, k\}$ and Γ is of full column rank k with $k < p(p+1)/2$, then the model is resistant to information switching.

Definition (Non-replicable Sparsity Pattern Condition)

All columns of Ψ^* , where Ψ^* is the $S_n \times k$ matrix stacking all the distinct ψ_s^* , are different.

Theorem

Let $\Gamma \in \mathbb{R}^{p \times k}$ be a real matrix with full column rank k . If $P \in \mathcal{O}_k$ and $\Gamma' = \Gamma P$ is a rotation of the specific factors, under the Non-replicable Sparsity Pattern Condition, then for each $s = 1, \dots, S_n$

$$\Lambda \Lambda^\top + \Gamma \Psi_s^* \Gamma^\top + \Sigma = \Lambda \Lambda^\top + \Gamma' \text{diag}(\psi_s^*) \Gamma'^\top + \Sigma,$$

if and only if P is a permutation matrix.

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Simulations: scenarios

We simulate data under the following scenarios

- *Scenario A* — *Correct specification*: $S = 3$ groups with sample size $n_1 = n_2 = n_3$, $d = 2$ active shared factors, and $k = 3$ ($[1 + 1 + 1]$ for the groups) specific factors
- *Scenario A** — *Latent Heterogeneity*: we do not provide the group labels but the data are generated as in Scenario A.
- *Scenario B* — *Homogeneity between groups*: While we provide $S = 3$ groups, the structure of latent factors is homogeneous among all the studies, i.e. $k = 0$.
- *Scenarios C and D* — *Mixed situations*: There exist groups but $k \neq S$



(A, A*)

(B)

(C)

(D)

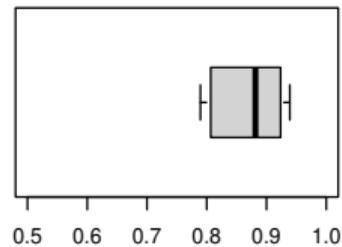
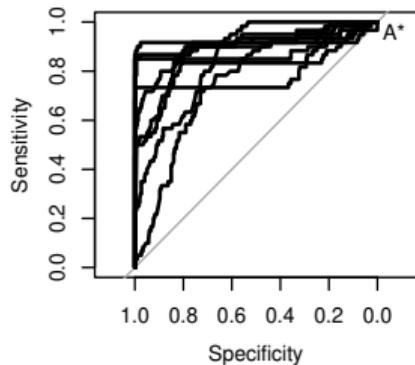
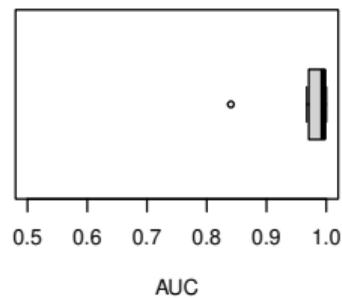
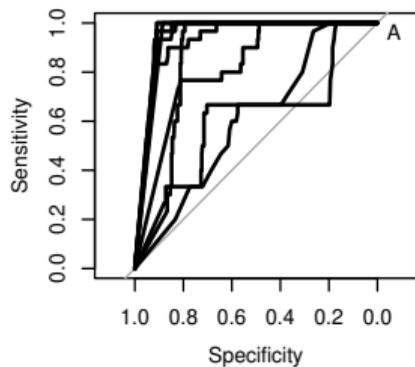
- We evaluate the performance of **APAFA** with the approach proposed Gabski et al., 2023 (**TETRIS**)
- For our method only, we evaluate the **ability in discovering the group structure**.
- To compare the relative performance of each competitor, we measure the adequacy of the reconstructed the variances of each group.

Results: variance matrices reconstruction

Table: Monte Carlo average (and interquartile range) of the posterior mean number of factors and RV coefficients for Ω_1 , Ω_2 , and Ω_3 on several simulation scenarios (the higher the better).

Scen	Method	d	k	Ω_1	Ω_2	Ω_3
A	APAFA	3.00 (1.00)	3.01 (0.23)	0.90 (0.07)	0.88 (0.07)	0.89 (0.05)
	TETRIS	3.67 (1.26)	4.91 (0.89)	0.92 (0.07)	0.93 (0.09)	0.88 (0.08)
A*	APAFA	4.00 (1.00)	3.00 (1.00)	0.69 (0.13)	0.78 (0.08)	0.75 (0.08)
B	APAFA	3.00 (0.00)	0.00 (0.00)	0.94 (0.04)	0.94 (0.04)	0.94 (0.04)
	TETRIS	3.00 (0.00)	0.00 (0.00)	0.90 (0.12)	0.92 (0.05)	0.92 (0.09)
C	APAFA	3.00 (0.00)	3.00 (0.05)	0.89 (0.05)	0.78 (0.04)	0.92 (0.02)
	TETRIS	3.00 (1.00)	2.00 (1.01)	0.72 (0.09)	0.75 (0.08)	0.79 (0.05)
D	APAFA	3.00 (0.00)	3.00 (0.12)	0.91 (0.03)	0.88 (0.03)	0.90 (0.04)

Results: group discovering





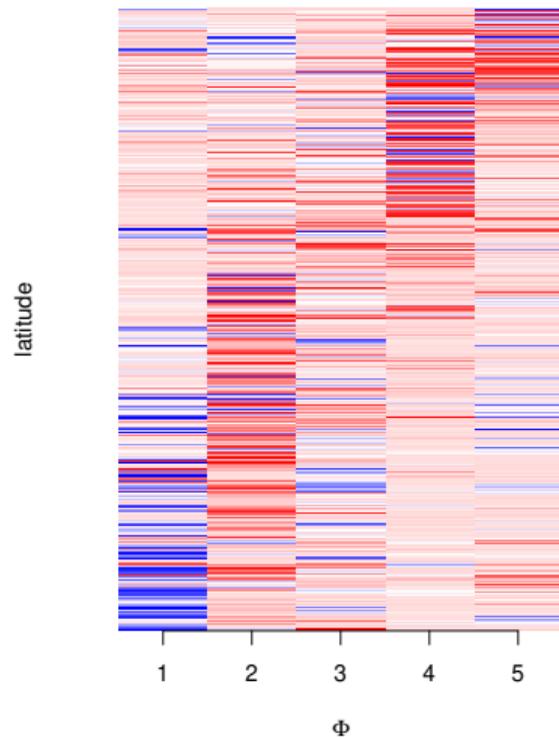
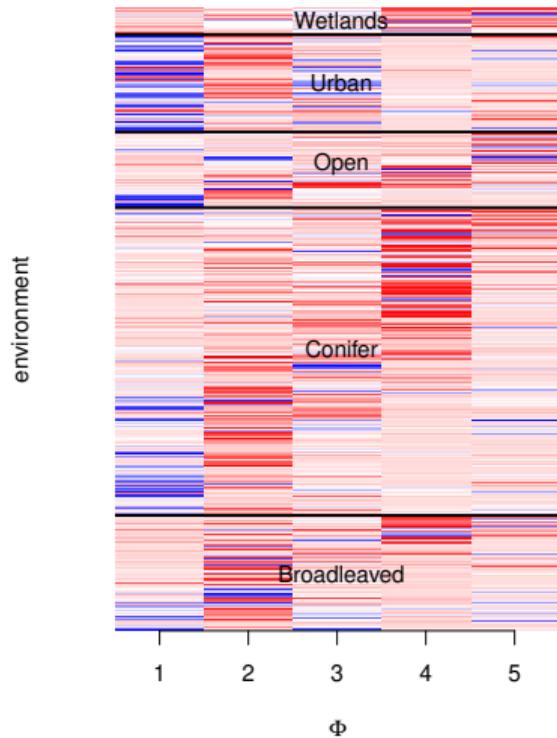
- we have a p -dimensional (binary) vector measuring the co-occurrence if p species of birds
- observations are grouped in $S = 200$ different locations, with $s = 1, \dots, S$.
- In each location we have repeated sampling campaigns for a total of n observation in total ($i = 1, \dots, n$)
- data y_{is} is the p -dimensional measurement for location s at sampling campaign i .

- The locations are considered as **groups in a multi-study framework**
- **y** : $n \times p$ binary matrix of occurrence of **p species** in **n different sampling campaigns**.
- We model species presence or absence using the multivariate probit regression model,

$$y_{ij} = 1(z_{ij} > 0), \quad z_i = \Lambda\eta_i + \Gamma\varphi_i + \epsilon_i, \quad \epsilon_i \sim N(0, \Sigma).$$

- **$S = 200$** : number of sites
- we do not use any information but the sampling campaign indicator. However ...
- The 200 locations can be clustered into 5 different types of location: Urban, Broadleaved forests, Coniferous forests, Open, and Wetlands.

Posterior estimate of specific factors



Wrapping up & essential references

- The concept of **sparsity** has been used to generalize the MSFA model
- The model not only enjoys appealing theoretical properties but shares many characteristics of **neural networks**
- APAFA exploits the broader concept of **structured shrinkage**:
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